

## A PROBABILISTIC APPROACH TO A BOUNDARY LAYER PROBLEM

BY

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**ABSTRACT.** An elliptic second order linear operator is approximated by the transition operator of a Markov chain, and the solution to the corresponding approximate boundary value problem is expanded in terms of a small parameter, up to the first order term. In characterizing the boundary values of the first order term in the expansion, a problem of a boundary layer arises, which is treated by probabilistic methods.

**1. Introduction.** Consider the boundary value problem:

$$(1.1) \quad \begin{aligned} Lu(x) &\equiv \sum_{i,j=1}^n a_{ij}(x) \frac{\partial^2}{\partial x_i \partial x_j} u(x) \\ &+ \sum_{j=1}^n b_j(x) \frac{\partial}{\partial x_j} u(x) = -g(x) \\ &\text{for } x \in G = \{x: \phi(x) < 0\}, x \in R^n, \end{aligned}$$

and

$$u(x) = b(x) \quad \text{for } x \in \partial G = \{x: \phi(x) = 0\},$$

where  $\phi(x) = 0$  is the boundary of a closed, bounded, connected domain  $G$  such that  $\nabla \phi(x) \neq 0$  for all  $x \in \partial G$ . Assume that  $a_{ij}(x)$ ,  $b_j(x)$ ,  $b(x)$ ,  $g(x)$ ,  $\phi(x)$  are in  $C_b^\infty$  and that the matrix  $(a_{ij}(x))^{n \times n}$  is uniformly positive definite. Let  $p_h(x, y)$  be a transition density of a Markov chain for which the jumps are uniformly bounded by  $ch^{1/2}$ , where  $c$  is a fixed positive constant and  $h$  a small parameter. We define a transition operator  $\Pi_h$  on  $C_b$  to be

$$(\Pi_h f)(x) = \int p_h(x, x') f(x') dx'.$$

Assume that

$$(1.2) \quad \lim_{h \rightarrow 0} \frac{(\Pi_h f)(x) - f(x)}{h} = Lf(x)$$

uniformly on compacts of  $R^n$  for each  $f \in C_b^2$ . We formulate the approximate problem of (1.1) to be

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$$(1.3) \quad \begin{aligned} (\Pi_h u_h)(x) - u_h(x) &= -hg(x), & x \in \bar{G}, \\ u_h(x) &= F(x), & x \in \bar{G}', \end{aligned}$$

where  $F(x)$  is a continuous extension of the boundary values  $b(x)$  into  $\bar{G}' = \{x: \phi(x) > 0\}$ . From the probabilistic point of view the solution to (1.3) is

$$u_h(x) = E_x^h F(x_{\tau_G}^h) + hE_x^h \left( \sum_{j=0}^{\tau_G-1} g(x_j^h) \right)$$

where  $\tau_G$  is the first exit time from  $\bar{G}$  of a Markov chain for which (1.2) holds. We obtain an asymptotic expansion of the form

$$(1.4) \quad u_h(x) = u(x) + \sqrt{h} u_1(x) + o(\sqrt{h})$$

where  $u(x)$  is the solution to (1.1) and  $u_1(x)$  is the solution to another boundary value problem, namely:

$$Lu_1(x) = -L_1 u(x), \quad x \in G;$$

where  $L_1$  is a differential operator determined by the Markov chain, and  $u_1(x) = b_1(x)$  for  $x \in \partial G$ , where  $b_1$  is to be determined. To see the kind of difficulties one can anticipate in characterizing  $b_1(x)$ , suppose we postulate that

$$(\Pi_h f)(x) = (I + hL + h^{3/2}L_1)(f)(x) + o(h^{3/2}),$$

$f \in C_b^\infty$ ; then the approximate equation (1.3) can be written in the form

$$(1.5) \quad (L + h^{1/2}L_1)u_h(x) = -g(x) + o(h^{1/2}).$$

Thus (1.3) can be viewed as an equation involving a singular perturbation of the operator  $L$ .

As an example of how the operator  $L_1$  is determined by the Markov chain, consider a probability distribution  $F(y)$  with zero mean, variance 1, third moment  $m_3 \neq 0$  and density  $f(x)$ ; then for  $a(x) \geq \delta > 0$ ,  $b(x)$  in  $C^\infty$

$$\frac{1}{\sqrt{ha(x)}} F \left[ \frac{y - x - hb(x)}{\sqrt{ha(x)}} \right]$$

is a distribution with mean  $x + hb(x)$  and variance 1. We make a change in variable,  $y = \sqrt{ha(x)} z + x + hb(x)$  and define  $\Pi_h$  as

$$\Pi_h u_h(x, y) = \int u_h(x, x + b(x)h + \sqrt{a(x)h} z) f(z) dz.$$

Now expanding formally by a Taylor expansion  $u_h(x, y)$  with respect to  $y$  at  $y = x$  and collecting terms in powers of  $\sqrt{h}$ , we obtain:

$$\begin{aligned} & \Pi_h u_h(x, y) \\ &= \left[ I + h \left( \frac{a(x)}{2} \frac{\partial^2}{\partial y^2} + b(x) \frac{\partial}{\partial y} \right) + h^{3/2} \left( \frac{m_3}{3!} (a(x))^{3/2} \frac{\partial^3}{\partial y^3} \right) \right] u_h(x, y) \\ & \quad + o(h^{3/2}). \end{aligned}$$

Thus

$$L_1 = \frac{m_3}{3!} (a(x))^{3/2} \frac{\partial^3}{\partial y^3}.$$

We know from singular perturbation theory that, in characterizing  $b_1(x)$ , we can anticipate a boundary layer problem. Employing probability methods, we treat the boundary layer problem and obtain the characterization of  $b_1(x)$ .

**2. The asymptotic expansion.** Since no complications arise in generalizing our results from dimension  $n = 2$  to  $n > 2$ , we shall limit our discussion to the two dimensional case. It is known that the  $\lim_{h \rightarrow 0} u_h(x) = u(x)$  for  $x \in G$ , see [4]. We begin the characterization of the first order term by proving the following:

**LEMMA 1.** *If*

$$\lim_{h \rightarrow 0} \frac{u_h(x) - u(x)}{h^{1/2}} = u_1(x)$$

*exists uniformly on compacts in  $G$ , then it is of the form  $u_1(x) = \bar{u}_1(x) + \hat{u}_1(x)$ , where  $\hat{u}_1(x)$  is the solution to  $L\hat{u}_1(x) = -L_1 u(x)$  with zero boundary conditions, and  $\bar{u}_1(x)$  is harmonic for  $L$  inside  $G$ .*

**PROOF.** Let  $U(x)$  be a smooth extension of  $u(x)$  into  $\bar{G}'$ , i.e., let the inward derivatives of  $u(x)$  agree with the corresponding outward derivatives of  $U(x)$  on  $\partial G$ . We note that  $u(x)$  possesses these inward derivatives since it is known that if the data are smooth then so is the solution in the case of elliptic boundary value problem [1]. It is easy to check that

$$U(x_n^h) - \sum_{j=0}^{\tau_G-1} (\Pi_h - I)U(x_j^h)$$

is a martingale, where  $x, x_1^h, x_2^h, x_n^h, \dots$  is a Markov chain satisfying the assumptions of §1. Substituting

$$(\Pi_h - I)U(x) = -hg(x) - h^{3/2}g_1(x) + o(h^{3/2}),$$

where  $L_1 u(x) = g_1(x)$ , in the martingale and taking into account that the expectation of a martingale is the same for all parameter values, that the expectation of the martingale at  $t = 0$  is  $U(x)$  and Doob's stopping theorem, we have

$$\begin{aligned}
E_x^h \left( U(x_{\tau_G}^h) - \sum_{j=0}^{\tau_G-1} (\Pi_h - I) U(x_{\tau_G}^h) \right) \\
= E_x^h U(x_{\tau_G}^h) + h E_x^h \sum_{j=0}^{\tau_G-1} g(x_j^h) \\
+ h^{3/2} E_x^h \sum_{j=0}^{\tau_G-1} g_1(x_j^h) + E_x^h(\tau_G) o(h^{3/2}) \\
= U(x).
\end{aligned}$$

Since  $E_x^h(\tau_G) \leq \text{const}/h$  for a discrete time Markov chain, see [2],

$$\frac{u_h(x) - U(x)}{\sqrt{h}} = \frac{1}{\sqrt{h}} \left( E_x^h (F(x_{\tau_G}^h) - U(x_{\tau_G}^h)) \right) - h E_x^h \sum_{j=0}^{\tau_G-1} g_1(x_j^h) + o(1).$$

But

$$\lim_{h \rightarrow 0} h E_x^h \left( \sum_{j=0}^{\tau_G-1} g_1(x_j^h) \right) = E_x \int_0^{\tau_G} g_1(x_t) dt,$$

where  $x_t$  is a diffusion process generated by the operator  $L$ . This limit (let us denote it by  $\hat{u}_1(x)$ ) is of course the solution to the Poisson equation  $L\hat{u}_1(x) = -g_1(x)$  with zero boundary conditions. Now let

$$\bar{u}_1^h(x) = \frac{1}{h^{1/2}} E_x^h (F(x_{\tau_G}^h) - U(x_{\tau_G}^h)).$$

By assumption  $\lim_{h \rightarrow 0} \bar{u}_1^h(x) = \bar{u}_1(x)$  exists for  $x \in G$ . Let  $D$  be a compact subset of  $G$ . Then by the strong Markov property

$$\bar{u}_1^h(x) = E_x^h \bar{u}_1^h(x_{\tau_D}^h) = \int \bar{u}_1^h(x) H_h^D(x, dy),$$

where  $\tau_D$  is the first exit time from  $D$  and  $H_h^D(x, dy)$  is the exit distribution from  $D$ . Since the  $\lim_{h \rightarrow 0} H_h^D(x, dy)$  is the harmonic measure on  $\partial D$ , it follows that  $\bar{u}_1(x)$  is harmonic in  $D$ . This completes the proof of Lemma 1.

It remains to prove that the hypothesis of Lemma 1 is satisfied and to obtain the boundary values of  $\bar{u}_1(x)$ . We begin by proving

LEMMA 2.

$$\lim_{\substack{h \rightarrow 0 \\ x \rightarrow b \in \partial G}} \left| \bar{u}_1^h(x) - W(b) \frac{1}{\sqrt{h}} E_x^h (\phi(x_{\tau_G}^h)) \right| = 0$$

where

$$W(x) = \frac{\langle \nabla (F(x) - U(x)), \nabla \phi(x) \rangle}{\| \nabla \phi(x) \|^2}$$

( $\langle \cdot, \cdot \rangle$  denotes inner-product).

PROOF. Let  $V(x) = F(x) - U(x)$ ; then

$$\bar{u}_1^h(x) = (1/h^{1/2})E_x^h V(x_{\tau_G}^h);$$

if we expand  $V(x)$  in the normal direction to  $\phi(x) = 0$ , we obtain

$$V(x) = \frac{\langle V(x), \nabla \phi(x) \rangle}{\|\nabla \phi(x)\|^2} \phi(x) + O(\phi^2(x)).$$

Since the jumps of the Markov chain are uniformly bounded by  $ch^{1/2}$ , we have  $O(\phi^2(x_{\tau_G}^h)) = O(h)$ . Thus

$$\bar{u}_1^h(x) = \frac{1}{\sqrt{h}} E_x^h (W(x_{\tau_G}^h) \phi(x_{\tau_G}^h)) + O(h^{1/2}).$$

Now for  $b$  on  $\partial G$

$$\begin{aligned} \lim_{\substack{h \rightarrow 0 \\ x \rightarrow b}} \frac{1}{\sqrt{h}} & \left| E_x^h (W(x_{\tau_G}^h) \phi(x_{\tau_G}^h)) - W(b) E_x^h \phi(x_{\tau_G}^h) \right| \\ &= \lim_{\substack{h \rightarrow 0 \\ x \rightarrow b}} \frac{1}{\sqrt{h}} \left| E_x^h [\phi(x_{\tau_G}^h) (W(x_{\tau_G}^h) - W(b))] \right| \\ &\leq \lim_{\substack{h \rightarrow 0 \\ x \rightarrow b}} k E_x^h |W(x_{\tau_G}^h) - W(b)| = 0. \end{aligned}$$

The last inequality follows from the assumption that the jumps of the Markov chain are uniformly bounded by  $ch^{1/2}$ . Assuming the continuity of  $W(x)$ , the last equality follows from the fact that  $V_n(x) = E_x^h [|W(x_{\tau_G}^h) - W(b)|]$  converges uniformly in  $x \in \bar{G}$  to the solution of  $LV = 0$  with  $V = |W(\cdot) - W(b)|$  on  $\partial G$ . This completes the proof of Lemma 2.

Thus, by Lemma 2, to identify the boundary value of  $\bar{u}_1(x)$  at  $b$ , we must identify

$$\lim_{x \rightarrow b} \lim_{h \rightarrow 0} \frac{1}{h^{1/2}} E_x^h \phi(x_{\tau_G}^h).$$

To this end we require several preliminary lemmas.

LEMMA 3. Assume the transition density  $p_h(x, y)$  is such that

$$\lim_{\substack{h \rightarrow 0 \\ x \rightarrow b}} h p_h(x, x + h^{1/2} z) = p_0^b(z),$$

where  $p_0^b(z)$  is a bivariate probability density. Then for  $f$  in  $C_b^\infty$

$$\lim_{\substack{h \rightarrow 0 \\ x \rightarrow b}} \iint_{R^2} \frac{f(\phi(y) - \phi(x))}{\sqrt{h}} p_h(x, y) dy = \int_R f(s) p^b(s) ds,$$

where  $p^b(s)$  is the marginal density of  $\langle \nabla \phi(b), Z \rangle$  and  $Z$  is a random vector with density  $p_0^b(z)$ .

PROOF. By a Taylor expansion

$$\phi(y) = \phi(x) + \langle \nabla \phi(x), (y - x) \rangle + O(\|y - x\|^2).$$

Thus

$$\begin{aligned} \iint f \left( \frac{\phi(y) - \phi(x)}{\sqrt{h}} \right) p_h(x, y) dy \\ = \iint f \left( \frac{\langle \nabla \phi(x), (y - x) \rangle + O(\|y - x\|^2)}{\sqrt{h}} \right) p_h(x, y) dy \\ = \iint \left( f \left( \frac{1}{\sqrt{h}} \langle \nabla \phi(x), (y - x) \rangle \right) + \frac{1}{\sqrt{h}} O(\|y - x\|^2) \right) p_h(x, y) dy \end{aligned}$$

where  $O(\|y - x\|^2) \leq k\|y - x\|^2$ . But assumption (1.2) implies

$$\lim_{h \rightarrow 0} \frac{1}{h} \iint \|y - x\|^2 p_h(x, y) dy = \text{trace}(a_{ij}(x));$$

thus

$$\lim_{h \rightarrow 0} \frac{1}{\sqrt{h}} \iint \|y - x\|^2 p_h(x, y) dy = 0.$$

Taking  $z = (y - x)/\sqrt{h}$  we have

$$\begin{aligned} \iint_{R^2} f \left( \frac{1}{\sqrt{h}} \langle \nabla \phi(x), (y - x) \rangle \right) p_h(x, y) dy \\ = \iint_{R^2} f(\langle \nabla \phi(x), z \rangle) h p_h(x, x + \sqrt{h} z) dz. \end{aligned}$$

But

$$\lim_{\substack{h \rightarrow 0 \\ x \rightarrow b}} h p_h(x, x + \sqrt{h} z) = p_0^b(z).$$

Therefore,

$$\lim_{\substack{h \rightarrow 0 \\ x \rightarrow b}} \iint_{R^2} f(\langle \nabla \phi(x), z \rangle) h p_h(x, x + \sqrt{h} z) dz = \int_R f(s) p^b(s) ds.$$

This completes the proof of Lemma 3.

LEMMA 4. Let

$$z_h(n) = (1/\sqrt{h}) \phi(x_n^h),$$

where  $\{x_n^h\}$  is a Markov chain starting at  $x \in G$  with transition density  $p_h(x, y)$ . Then for every  $k$ ,  $z_h(1) - z_h(0)$ ,  $z_h(2) - z_h(1)$ ,  $\dots$ ,  $z_h(k) - z_h(k-1)$  converge weakly to  $z(1) - z(0)$ ,  $z(2) - z(1)$ ,  $\dots$ ,  $z(k) - z(k-1)$  as  $h \rightarrow 0$  and  $x \rightarrow b$  where  $z(n)$  ( $n = 0, 1, 2, \dots$ ) is a random walk with transition density  $p^b(s)$ .

PROOF. To show that any finite number of increments of  $z_h(n)$  ( $n = 0, 1, 2, \dots$ ) converge weakly to the corresponding finite number of increments of  $z(n)$  ( $n = 0, 1, 2, \dots$ ), it suffices to show that for any  $k$ ,

$$\lim_{\substack{h \rightarrow 0 \\ x \rightarrow b}} E_x^h (f_1(z_h(1) - z_h(0)) f_2(z_h(2) - z_h(1)) \cdots f_k(z_h(k) - z_h(k-1))) \\ = E_b (f_1(z(1) - z(0)) f_2(z(2) - z(1)) \cdots f_k(z(k) - z(k-1)))$$

where  $f_j$  ( $j = 1, 2, \dots, k$ ) are any continuous, bounded functions. Now

$$E_x^h f_1(z_h(1) - z_h(0)) f_2(z_h(2) - z_h(1)) \cdots f_k(z_h(k) - z_h(k-1)) \\ = \int_{R^{2k}} f_1 \left( \frac{\phi(x_1) - \phi(x)}{\sqrt{h}} \right) f_2 \left( \frac{\phi(x_2) - \phi(x_1)}{\sqrt{h}} \right) \cdots f_k \left( \frac{\phi(x_k) - \phi(x_{k-1})}{\sqrt{h}} \right) \\ \times p_h(x, x_1) p_h(x_1, x_2) \cdots p_h(x_{k-1}, x_k) dx_1 dx_2 \cdots dx_k.$$

We shall evaluate one factor at a time; in the limit this will permit us to write the above expectation as a product of expectations. Evaluating the limit of the  $k$ th factor first, i.e., by Lemma 3

$$\lim_{\substack{h \rightarrow 0 \\ x_{k-1} \rightarrow b}} \int \int_{R^2} f_k \left( \frac{\phi(x_k) - \phi(x_{k-1})}{\sqrt{h}} \right) dx_k = \int_R f(s) p^b(s) ds, \quad x_k \in R^2.$$

Similarly, by Lemma 3 we evaluate the limit of the  $(k-1)$ th factor, and again obtain  $\int_R f(s) p^b(s) ds$ . Continuing to evaluate each consecutive factor in the same manner, we obtain

$$\lim_{\substack{h \rightarrow 0 \\ x_i \rightarrow b \\ i = k-1, k-2, \dots, 0}} E_x^h f_1(z_h(1) - z_h(0)) \cdots f_k(z_h(k) - z_h(k-1)) \\ = \int_R f_1(s) p^b(s) ds \cdots \int_R f_k(s) p^b(s) ds.$$

This implies that  $z(i) - z(i-1)$  ( $i = 1, 2, \dots, k$ ) are independent and identically distributed, and thus  $z(n)$ ,  $n = 0, 1, 2, \dots$ , is a random walk. This completes the proof of Lemma 4.

LEMMA 5.

$$\lim_{\substack{h \rightarrow 0 \\ x \rightarrow b \\ \phi(x)/\sqrt{h} \rightarrow -s}} \frac{1}{\sqrt{h}} E_x^h (\phi(x_{\tau_G}^h)) = E_{-s}^b (z(\tau_G))$$

where  $x_n^h$  is a Markov chain with transition density  $p_h(x, y)$  and  $z(n)$  is a random walk with transition density  $p^b(s)$ .

REMARK. The

$$\lim_{h \rightarrow 0} \frac{1}{\sqrt{h}} E_x^h(\phi(x_{\tau_G}^h))$$

is nonuniform near the boundary of  $G$ ; thus great care must be taken in how we let  $h \rightarrow 0$  and  $x \rightarrow b$ . Above we let  $h \rightarrow 0$  and  $x \rightarrow b$  in such a way that

$$\phi(x)/\sqrt{h} \rightarrow -s$$

then in order to evaluate

$$\lim_{x \rightarrow b} \lim_{h \rightarrow 0} \frac{1}{\sqrt{h}} E_x^h(\phi(x_{\tau_G}^h))$$

we shall let  $s \rightarrow \infty$ .  $E_{-s}^b(z(\tau_G))$  is the expected overshoot in the normal direction over the boundary at  $x = b$  of the random walk  $z(n)$  starting at  $-s$ .

PROOF. We note that

$$z_h(\tau_G) = (1/\sqrt{h})\phi(x_{\tau_G}^h);$$

thus

$$z_h(0) = \phi(x)/\sqrt{h},$$

and thus

$$(1/\sqrt{h})E_x^h(\phi(x_{\tau_G}^h)) = E_{\phi(x)/\sqrt{h}}(z_h(\tau_G)).$$

By Lemma 4 the finite dimensional distributions of the stochastic process  $z_h(n)$  ( $h = 0, 1, 2, \dots$ ) converge weakly to the finite dimensional distributions of the random walk  $z(n)$  ( $n = 0, 1, 2, \dots$ ). Since  $\tau_G = \inf_n \{n: z_h(n) > 0\}$  is finite almost surely, it follows by a standard weak convergence argument that

$$\lim_{\substack{h \rightarrow 0 \\ x \rightarrow b \\ \phi(x)/\sqrt{h} \rightarrow -s}} E_{\phi(x)/\sqrt{h}}^h z_h(\tau_G) = E_{-s}^b(z(\tau_G)).$$

This completes the proof of Lemma 5.

REMARK. Assuming that  $p_0^b(z)$  in Lemma 3 is such that  $\iint z p_0^b(z) dz = 0$ , then the random walk  $z(n)$  will be persistent and by the Renewal Theorem  $\lim_{s \rightarrow \infty} E_{-s}^b z(\tau_G)$  exists; it is shown in [3] that for a persistent, nonlattice random walk the limiting exit distribution exists.

We shall assume:

$$(2.1) \quad \lim_{s \rightarrow \infty} E_{-s}^b z(\tau_G) \text{ exists uniformly in } b \in \partial G.$$

In (2.1) we are assuming that if two boundary points are close then so are the expected overshoots, in the normal directions to the boundary at these points, of the random walk  $z(n)$ . For this it is enough to assume sufficient smoothness of the boundary and the transition density.

Denote  $E_{-s}^b(z(\tau_G))$  by  $V_s(b)$  and  $\lim_{s \rightarrow \infty} V_s(b)$  by  $V(b)$ . Our next objective



is to show that  $\lim_{h \rightarrow 0} \bar{u}_1^h(x) = \bar{u}_1(x)$  exists,  $L\bar{u}_1 = 0$ , and  $\bar{u}_1(x)$  takes on the boundary value  $W(b)V(b)$  at  $b \in \partial G$ . To this end we first prove the following

LEMMA 6. Let  $P_n$  be a sequence of probability measures on a measurable space  $(X, \mathcal{F})$  converging weakly to a probability measure  $P$ , where  $X$  is a complete separable metric space and  $\mathcal{F}$  its Borel field. Let  $C$  be a closed set in  $\mathcal{F}$  such that  $P(C) = 1$ ; let  $f_n(x)$  be a sequence of measurable functions on  $X$  and  $f(x)$  a measurable function on  $X$ . Furthermore assume

(1)  $|f_n(x)| \leq K$  for all  $n$ , and

(2)  $\lim_{n \rightarrow \infty; x_n \rightarrow x \in C} |f_n(x_n) - f(x)| \leq \delta$ ;

then

$$\overline{\lim}_{n \rightarrow \infty} \left| \int_X f_n(x) dP_n(x) - \int_X f(x) dP(x) \right| \leq \delta.$$

PROOF. Suppose the conclusion is not true; then there exist a subsequence  $f_{n_j}(x)$  and  $P_{n_j}$  such that

$$\overline{\lim}_{n_j} \left| \int_X f_{n_j}(x) dP_{n_j} - \int_X f(x) dP(x) \right| = \delta' > \delta.$$

But  $P_{n_j}$  converges weakly to  $P$ ; thus by a theorem of Skorohod there exist a probability space  $(\Omega, \Sigma, \Pi)$  and random variables  $x_{n_j}(\omega)$ ,  $x(\omega)$  such that  $\Pi(x_{n_j}^{-1}(A)) = P_{n_j}(A)$ , and  $\Pi(x^{-1}(A)) = P(A)$  for all  $A \in \mathcal{F}$ ; further,  $x_{n_j}(\omega)$  converges in probability to  $x(\omega)$ , and there exists a further subsequence  $x_{n_{j_k}}(\omega)$  which converges to  $x(\omega)$ , almost surely. Thus by assumption (2)

$$\overline{\lim}_{n \rightarrow \infty} |f_{n_{j_k}}(x_{n_{j_k}}(\omega)) - f(x(\omega))| \leq \delta$$

a.e., and by the Bounded Convergence Theorem

$$\overline{\lim}_{n_{j_k}} \left| \int_{\Omega} f_{n_{j_k}}(x_{n_{j_k}}(\omega)) d\Pi(\omega) - \int_{\Omega} f(x(\omega)) d\Pi(\omega) \right| \leq \delta.$$

But this contradicts the assumption that

$$\overline{\lim}_{n_j} \left| \int_X f_{n_j}(x) dP_{n_j}(x) - \int_X f(x) dP(x) \right| = \delta' > \delta.$$

This completes the proof of Lemma 6.

LEMMA 7.  $\lim_{h \rightarrow 0} \bar{u}_1^h(x) = \bar{u}_1(x)$  exists;  $L\bar{u}_1(x) = 0$  for  $x \in G$  and  $\bar{u}_1(b) = W(b)V(b)$ ,  $b \in \partial G$ .

PROOF. Let  $\tau_{s,h}$  be the first exit time of a Markov chain  $\{x_n^h\}$  with transition density  $p_h(x, y)$  from  $\{x: \phi(x) < -sh^{1/2}\}$ ,  $s > k$ , where  $kh^{1/2}$  is the uniform bound on the jumps of  $\phi(x_n^h)$ . Then by assumption (2.1), Lemma 2 and Lemma 5 we have

$$\lim_{\substack{h \rightarrow 0 \\ x \rightarrow b \\ \phi(x)/\sqrt{h} \rightarrow s'}} |\bar{u}_1^h(x) - W(b)V_{s'}(b)| = 0,$$

where  $V_{s'}(b) = E_{-s'}^b(z(\tau_G))$ . From this it follows that for every  $b \in \partial G$ ,

$$\begin{aligned} \lim_{\substack{h \rightarrow 0 \\ x \rightarrow b \\ -s\sqrt{h} < \phi(x) < -s\sqrt{h} + c\sqrt{h}}} |\bar{u}_1^h(x) - W(b)V(b)| \\ &< \sup_{s-c < s' < s} |W(b)V_{s'}(b) - W(b)V(b)| \\ &< \sup_b |W(b)| \sup_{s-c < s' < s} \sup_b |V_{s'}(b) - V(b)| \\ &= \sup_b |W(b)| \mathcal{E}(s-c) \end{aligned}$$

where  $\mathcal{E}(s) = \sup_{s' > s} \sup_b |V_{s'}(b) - V(b)|$ . By our assumption (2.1)  $\mathcal{E}(s) \rightarrow 0$  as  $s \rightarrow \infty$ . Now by the strong Markov property

$$\bar{u}_1^h(x) = E_x^h \bar{u}_1(x_{\tau_h}^h) = \int \int_D \bar{u}_1^h(x) H_h^s(x, dy),$$

$$D = \{x: -sh^{1/2} < \phi(x) < -sh^{1/2} + ch^{1/2}\},$$

where  $H_h^s(x, dy)$  is the exit distribution from  $\{s: -sh^{1/2} > \phi(x)\}$ . It is well known that for regular boundary points  $b \in \partial G$ ,

$$\lim_{\substack{h \rightarrow 0 \\ x \rightarrow b}} H_h^s(x, dy) = H(x, db),$$

where  $H(x, db)$  is the harmonic measure on  $\partial G$ . Setting  $C = \partial G$ ,  $f_n = \bar{u}_1^h$ ,  $f = \bar{u}$ ,  $P_n = H_h^s$ ,  $P = H$  and  $\mathcal{E}(s) = \delta$  in Lemma 6, and noting that all the assumptions of Lemma 6 are met, we can conclude by Lemma 6 that

$$\begin{aligned} \overline{\lim}_{h \rightarrow 0} |\bar{u}_1^h(x) - \bar{u}_1(x)| \\ &= \overline{\lim}_{h \rightarrow 0} \left| \int \int_D \bar{u}_1^h(x) H_h^s(x, dy) - \int_{\partial G} W(b)V(b)H(x, db) \right| \\ &< \mathcal{E}(s-c). \end{aligned}$$

But  $\bar{u}_1^h(x)$  and  $\bar{u}_1(x)$  are independent of  $s$ , and  $\mathcal{E}(s-c) \rightarrow 0$  as  $s \rightarrow \infty$ . This completes the proof of Lemma 7.

Recalling that  $u_1(x) = \hat{u}_1(x) + \bar{u}_1(x)$  where  $\hat{u}_1(x)$  satisfies the Poisson equation,  $L\hat{u}_1 = -g_1$ , with zero boundary condition and  $\bar{u}_1(x)$  is harmonic for  $L$  inside  $G$  with the boundary values given by Lemma 7. We can now summarize our results by the following

**THEOREM.** *Let*

$$A = \{x: \phi(x) \leq d < 0\},$$

$$V(b) = \lim_{\substack{h \rightarrow 0 \\ x \rightarrow b \\ \phi(x)/\sqrt{h} \rightarrow \infty}} \frac{1}{\sqrt{h}} E_x^h \phi(x_{\tau_G}^h)$$

and

$$W(x) = \frac{\langle \nabla (F(x) - U(x)), \nabla \phi(x) \rangle}{\|\nabla \phi(x)\|^2},$$

then

$$\lim_{h \rightarrow 0} \sup_{x \in A} \left| \frac{u_h(x) - u(x)}{\sqrt{h}} - u_1(x) \right| = 0,$$

where  $u_1(x)$  is the solution to  $Lu_1(x) = -g_1(x)$  with boundary values  $u_1(b) = W(b)V(b)$ ,  $b \in \partial G$ .

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#### REFERENCES

1. L. Bers, F. John and M. Schecter, *Partial differential equations*, Lectures in Appl. Math., Vol. III, Wiley, New York, 1964. MR 29 #346.
2. E. Dynkin, *Infinitesimal operators of Markov processes*, English transl., Theor. Probability Appl. 1 (1956), 34-54. MR 19, 691.
3. W. Feller, *An introduction to probability theory and its applications*, Vol. 2, 2nd ed., Wiley, New York, 1971. MR 42 #5292.
4. G. Forsythe and W. Wasow, *Finite-difference methods for partial differential equations*, Wiley, New York, 1960. MR 23 #B3156.

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